Moments of k-hop counts in the random-connection model

Nicolas Privault*

Division of Mathematical Sciences School of Physical and Mathematical Sciences Nanyang Technological University 21 Nanyang Link Singapore 637371

July 3, 2019

Abstract

We derive moment identities for the stochastic integrals of multiparameter processes in a random-connection model based on a point process admitting a Papangelou intensity. Those identities are written using sums over partitions, and they reduce to sums over non-flat partition diagrams in case the multiparameter processes vanish on diagonals. As an application, we obtain general identities for the moments of k-hop counts in the random-connection model, which simplify the derivations available in the literature.

Key words: Point processes, moments, random-connection model, random graph, *k*-hops.

Mathematics Subject Classification (2010): 60G57; 60G55.

1 Introduction

The random-connection model, see e.g. Chapter 6 of [12], is a classical model in continuum percolation. It consists in a random graph built on the vertices of a point process on \mathbb{R}^d , by adding edges between two distinct vertices x and y with probability H(||x - y||). In the case of the Rayleigh fading $H_{\beta}(||x - y||) = e^{-\beta ||x - y||^2}$ with $x, y \in \mathbb{R}^2$, the mean value of the number $N_k^{x,y}$ of k-hop paths connecting $x \in \mathbb{R}^d$

^{*}nprivault@ntu.edu.sg

to $y \in \mathbb{R}^d$ has been computed in [9], together with the variance of 3-hop counts. However, this argument does not extend to $k \geq 3$ as the proof of the variance identity for 3-hop counts in [9] relies on the known Poisson distribution of the 2-hop count. As shown by [9], the knowledge of moments can provide accurate numerical estimates of the probability $P(N_k^{x,y} > 0)$ of at least one k-hop path, by expressing it as a series of factorial moments, and the need for a general theory of such moments has been pointed out therein.

On the other hand, moment identities for Poisson stochastic integrals with random integrands have been obtained in [18] based on moment identities for Skorohod's integral on the Poisson space, see [16, 17], and also [19] for a review. These moment identities have been extended to point processes with Papangelou intensities by [5], and to multiparameter processes by [2]. Factorial moments have also been computed by [4] for point processes with Papangelou intensities.

In this paper we derive closed-form expressions for the moments of the number of k-hop paths in the random-connection model. In Proposition 3.1 the moment of order n of the k-hop count is given as a sum over non-flat partitions of $\{1, \ldots, nk\}$ in a general random-connection model based on a point process admitting a Papangelou intensity. Those results are then specialized to the case of Poisson point processes, with an expression for the variance of the k-hop count given in Corollary 3.2 using a sum over integer sequences. Finally, in the case of Rayleigh fadings we show that some results of [9], such as the computation of variance for 3-hop counts, can be recovered via a shorter argument, see Corollary 5.3.

We proceed as follows. After presenting some background notation on point processes and Campbell measures, see [8], in Section 2 we review the derivation of moment identities for stochastic integrals using sums over partitions. In the multiparameter case we rewrite those identities for processes vanishing on diagonals, based on nonflat partition diagrams. In Section 3 we apply those results to the computation of the moments of k-hop counts in the random-connection model, and we specialize such computations to the case of Poisson point processes in Section 4. Section 5 is devoted to explicit computations in the case of Rayleigh fadings, which result into simpler derivations in comparison with the current literature on moments in the randomconnection model.

Notation on point processes

Let X be a Polish space with Borel σ -algebra $\mathcal{B}(X)$, equipped with a σ -finite nonatomic measure $\lambda(dx)$. We let

$$\Omega^X = \left\{ \omega = \{x_i\}_{i \in I} \subset X : \#(A \cap \omega) < \infty \text{ for all compact } A \in \mathcal{B}(X) \right\}$$

denote the space of locally finite configurations on X, whose elements $\omega \in \Omega^X$ are identified with the Radon point measures $\omega = \sum_{x \in \omega} \epsilon_x$, where ϵ_x denotes the Dirac measure at $x \in X$. A point process is a probability measure P on Ω^X equipped with the σ -algebra \mathcal{F} generated by the topology of vague convergence.

Point processes can be characterized by their Campbell measure C defined on $\mathcal{B}(X) \otimes \mathcal{F}$ by

$$C(A \times B) := \mathbb{E}\left[\int_{A} \mathbb{1}_{B}(\omega \setminus \{x\}) \ \omega(dx)\right], \quad A \in \mathcal{B}(X), \quad B \in \mathcal{F},$$

which satisfies the Georgii-Nguyen-Zessin [14] identity

$$\mathbb{E}\left[\int_{X} u(x;\omega)\omega(dx)\right] = \mathbb{E}\left[\int_{\Omega^{X}} \int_{X} u(x;\omega\cup x)C(dx,d\omega)\right],$$
(1.1)

for all measurable processes $u: X \times \Omega^X \to \mathbb{R}$ such that both sides of (1.1) make sense.

In the sequel we deal with point processes whose Campbell measure $C(dx, d\omega)$ is absolutely continuous with respect to $\lambda \otimes P$, i.e.

$$C(dx, d\omega) = c(x; \omega)\lambda(dx)P(d\omega),$$

where the density $c(x; \omega)$ is called the Papangelou density. We will also use the random measure $\hat{\lambda}^n(d\mathfrak{x}_n)$ defined on X^n by

$$\hat{\lambda}^n(d\mathfrak{x}_n) = \hat{c}(\mathfrak{x}_n;\omega)\lambda(dx_1)\cdots\lambda(dx_n),$$

where $\hat{c}(\mathfrak{x}_n;\omega)$ is the compound Campbell density $\hat{c}: \Omega_0^X \times \Omega^X \longrightarrow \mathbb{R}_+$ defined inductively on the set Ω_0^X of finite configurations in Ω^X by

$$\hat{c}(\{x_1, \dots, x_n, y\}; \omega) := c(y; \omega)\hat{c}(\{x_1, \dots, x_n\}; \omega \cup \{y\}), \qquad n \ge 0, \qquad (1.2)$$

see Relation (1) in [5]. In particular, the Poisson point process with intensity $\lambda(dx)$ is a point process with Campbell measure $C = \lambda \otimes P$ and $c(x;\omega) = 1$, and in this case the identity (1.1) becomes the Slivnyak-Mecke formula [20], [11]. Determinantal point processes are examples of point processes with Papangelou intensities, see e.g. Theorem 2.6 in [6], and they can be used for the modeling of wireless networks with repulsion, see e.g. [7], [13], [10].

2 Moment identities

The moment of order $n \ge 1$ of a Poisson random variable Z_{α} with parameter $\alpha > 0$ is given by

$$\mathbb{E}[Z_{\alpha}^{n}] = \sum_{k=0}^{n} \alpha^{k} S(n,k), \qquad n \in \mathbb{N},$$
(2.1)

where the Stirling number of the second kind S(n, k) is the number of ways to partition a set of n objects into k non-empty subsets, see e.g. Proposition 3.1 of [3]. Regarding Poisson stochastic integrals of deterministic integrands, in [1] the moment formula

$$\mathbb{E}\left[\left(\int_X f(x)\omega(dx)\right)^n\right] = n! \sum_{\substack{r_1+2r_2+\dots+nr_n=n\\r_1,\dots,r_n \ge 0}} \prod_{k=1}^n \left(\frac{1}{(k!)^{r_k}r_k!} \left(\int_X f^k(x)\lambda(dx)\right)^{r_k}\right) \quad (2.2)$$

has been proved for deterministic functions $f \in \bigcap_{p \ge 1} L^p(X, \lambda)$.

The identity (2.2) has been rewritten in the language of sums over partitions, and extended to Poisson stochastic integrals of random integrands in Proposition 3.1 of [18], and further extended to point processes admitting a Panpangelou intensity in Theorem 3.1 of [5], see also [4]. In the sequel, given $\mathfrak{z}_n = (z_1, \ldots, z_n) \in X^n$, we will use the shorthand notation $\varepsilon_{\mathfrak{z}_n}^+$ for the operator

$$(\varepsilon_{j_n}^+ F)(\omega) = F(\omega \cup \{z_1, \dots, z_n\}), \qquad \omega \in \Omega,$$

where F is any random variable on Ω^X . Given $\rho = \{\rho_1, \ldots, \rho_k\} \in \Pi[n]$ a partition of $\{1, \ldots, n\}$ of size $|\rho| = k$, we let $|\rho_i|$ denote the cardinality of each block ρ_i , $i = 1, \ldots, k$.

Proposition 2.1 Let $u: X \times \Omega^X \longrightarrow \mathbb{R}$ be a (measurable) process. For all $n \ge 1$ we have

$$\mathbb{E}\left[\left(\int_X u(x;\omega)\omega(dx)\right)^n\right] = \sum_{\rho\in\Pi[n]} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon^+_{\mathfrak{z}_{|\rho|}} \prod_{l=1}^{|\rho|} u^{|\rho_l|}(z_l)\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right],$$

where the sum runs over all partitions ρ of $\{1, \ldots, n\}$ with cardinality $|\rho|$.

Proposition 2.1 has also been extended, together with joint moment identities, to multiparameter processes $(u_{z_1,\ldots,z_r})_{(z_1,\ldots,z_r)\in X^r}$, see Theorem 3.1 of [2]. For this, let $\Pi[n \times r]$ denote the set of all partitions of the set

$$\Delta_{n \times r} := \{1, \dots, n\} \times \{1, \dots, r\} = \{(k, l) : k = 1, \dots, n, l = 1, \dots, r\},\$$

identified to $\{1, \ldots, nr\}$, and let $\pi := (\pi_1, \ldots, \pi_n) \in \Pi[n \times r]$ denote the partition made of the *n* blocks $\pi_k := \{(k, 1), \ldots, (k, r)\}$ of size *r*, for $k = 1, \ldots, n$. Given $\rho = \{\rho_1, \ldots, \rho_m\}$ a partition of $\Delta_{n \times r}$, we let $\zeta^{\rho} : \Delta_{n \times r} \longrightarrow \{1, \ldots, m\}$ denote the mapping defined as

$$\zeta^{\rho}(k,l) = p \text{ if and only if } (k,l) \in \rho_p, \qquad (2.3)$$

k = 1, ..., n, l = 1, ..., r, p = 1, ..., m. In other words, $\zeta^{\rho}(k, l)$ denotes the index p of the block $\rho_p \subset \Delta_{n \times r}$ to which (k, l) belongs.

Next, we restate Theorem 3.1 of [2] by noting that, in the same way as in Proposition 2.1, it extends to point processes admitting a Papangelou intensity using the arguments of [5], [4]. When $(u(z_1, \ldots, z_k; \omega))_{z_1, \ldots, z_k \in X}$ is a multiparameter process, we will write

$$\epsilon_{\mathfrak{z}_k}^+ u(z_1,\ldots,z_k;\omega) := u\big(z_1,\ldots,z_k;\omega \cup \{z_1,\ldots,z_k\}\big), \quad \mathfrak{z}_n = (z_1,\ldots,z_n) \in X^n,$$

and in this case we may drop the variable $\omega \in \Omega^X$ by writing $\epsilon_{\mathfrak{z}_k}^+ u(z_1, \ldots, z_k; \omega)$ instead of $\epsilon_{\mathfrak{z}_k}^+ u(z_1, \ldots, z_k; \omega)$.

Proposition 2.2 Let $u: X^r \times \Omega^X \longrightarrow \mathbb{R}$ be a (measurable) r-process. We have

$$\mathbb{E}\left[\left(\int_{X^r} u(z_1, \dots, z_r; \omega)\omega(dz_1) \cdots \omega(dz_r)\right)^n\right] = \sum_{\rho \in \Pi[n \times r]} \mathbb{E}\left[\int_{X^{|\rho|}} \varepsilon_{\mathfrak{z}_{|\rho|}}^+ \prod_{k=1}^n u(z_{\pi_k}^{\rho})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right]$$
(2.4)

with $z_{\pi_k}^{\rho} := (z_{\zeta^{\rho}(k,1)}, \dots, z_{\zeta^{\rho}(k,r)}), \ k = 1, \dots, n.$

Proof. The main change in the proof argument of [2] is to rewrite the proof of Lemma 2.1 therein by applying (1.2) recursively as in the proof of Theorem 3.1 of [5], while the main combinatorial argument remains identical.

When n = 1, Proposition 2.2 yields a multivariate version of the Georgii-Nguyen-Zessin identity (1.1), i.e.

$$\mathbb{E}\left[\int_{X^r} u(z_1,\ldots,z_r;\omega)\omega(dz_1)\cdots\omega(dz_r)\right] = \sum_{\rho\in\Pi[1\times r]} \mathbb{E}\left[\int_{X^{|\rho|}} \varepsilon_{\mathfrak{z}_{|\rho|}}^+ u(z_{\zeta^{\rho}(1,1)},\ldots,z_{\zeta^{\rho}(1,r)};\omega)\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right].$$

Non-flat partitions

In the sequel we write $\nu \leq \sigma$ when a partition $\nu \in \Pi[n \times r]$ is finer than another partition $\sigma \in \Pi[n \times r]$, i.e. when every block of ν is contained in a block of σ , and we let $\hat{0} := \{\{1, 1\}, \ldots, \{n, r\}\}$ denote the partition of $\Delta_{n \times r}$ made of singletons. We write $\rho \wedge \nu = \hat{0}$ when $\mu = \hat{0}$ is the only partition $\mu \in \Pi[n \times r]$ such that $\mu \leq \nu$ and $\mu \leq \rho$, i.e. $|\nu_k \cap \rho_l| \leq 1, \ k = 1, \ldots, n, \ l = 1, \ldots, |\rho|$. In this case we say that the partition diagram $\Gamma(\nu, \rho)$ of ν and ρ is *non-flat*, see Chapter 4 of [15].

In the sequel, a partition $\rho \in \Pi[n \times r]$ is said to be *non-flat* if the partition diagram $\Gamma(\pi, \rho)$ of ρ and the partition π is *non-flat*, where $\pi := (\pi_1, \ldots, \pi_n) \in \Pi[n \times r]$ with $\pi_k := \{(k, 1), \ldots, (k, r)\}, k = 1, \ldots, n$. The following figure shows an example of a non-flat partition

π_5	×	\triangle	۲	
π_4		×		۲
π_3	۲			×
π_2	\bigtriangleup	٢	×	
π_1	۲	\triangle		٢

with n = 5, r = 4, and

$$\begin{split} & \triangle = \{(1,2), (2,1), (2,2), (3,3), (4,2)\}, \\ & \bigcirc = \{(1,1), (3,1), (4,4), (5,3)\}, \\ & \square = \{(1,3), (2,4), (3,3), (4,1), (5,4)\}, \\ & \triangle = \{(1,4), (2,2)\}, \\ & \times = \{(2,3), (3,4), (4,2), (5,1)\} \\ & \pi_k = \{(k,1), (k,2), (k,3), (k,4), (k,5)\}, \quad k = 1, 2, 3, 4, 5. \end{split}$$

Processes vanishing on diagonals

The next consequence of Proposition 2.2 shows that when $u(z_1, \ldots, z_r; \omega)$ vanishes on the diagonals in X^r , the moments of $\int_{X^r} u(z_1, \ldots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)$ reduce to sums over non-flat partition diagrams.

Proposition 2.3 Assume that $u(z_1, \ldots, z_r; \omega) = 0$ whenever $z_i = z_j, 1 \le i \ne j \le r$, $\omega \in \Omega^X$. Then we have

$$\mathbb{E}\left[\left(\int_{X^r} u(z_1,\ldots,z_r;\omega)\omega(dz_1)\cdots\omega(dz_r)\right)^n\right] = \sum_{\substack{\rho\in\Pi[n\times r]\\\rho\wedge\pi=\hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon^+_{\mathfrak{z}_{|\rho|}} \prod_{k=1}^n u(z^{\rho}_{\pi_k})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right]$$

Proof. Assume that $u(z_1, \ldots, z_r; \omega)$ vanishes on diagonals, and let $\rho \in \Pi[n]$. Then, for any $z_1, \ldots, z_r \in X$ we have

$$\prod_{k=1}^{n} u(z_{\pi_{k}}^{\rho}) = \prod_{k=1}^{n} u(z_{\zeta^{\rho}(k,1)}, \dots, z_{\zeta^{\rho}(k,r)}) = 0$$

whenever $p := \zeta^{\rho}(k, a) = \zeta^{\rho}(k, b)$ for some $k \in \{1, \ldots, n\}$ and $a \neq b \in \{1, \ldots, r\}$. According to (2.3) this implies $(k, a) \in \rho_p$ and $(k, b) \in \rho_p$, therefore ρ is not a non-flat partition, and it should be excluded from the sum over $\Pi[n]$.

When n = 1, the first moment in Proposition 2.3 yields the Georgii-Nguyen-Zessin identity

$$\mathbb{E}\left[\int_{X^r} u(z_1,\ldots,z_r;\omega)\omega(dz_1)\cdots\omega(dz_r)\right] = \sum_{\substack{\rho\in\Pi[1\times r]\\\rho\wedge\pi=\hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon^+_{\mathfrak{z}_{|\rho|}} u(z^{\rho}_{\pi_1})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right]$$

$$= \mathbb{E}\left[\int_{X^r} \epsilon_{\mathfrak{z}_r}^+ u(z_1, \dots, z_r; \omega) \hat{\lambda}^r(d\mathfrak{z}_r)\right] (2.5)$$

see Lemma IV.1 in [9] and Lemma 2.1 in [2] for different versions based on the Poisson point process. In the case of second moments, we find

$$\mathbb{E}\left[\left(\int_{X^r} u(z_1,\ldots,z_r;\omega)\omega(dz_1)\cdots\omega(dz_r)\right)^2\right] = \sum_{\substack{\rho\in\Pi[2\times r]\\\rho\wedge\pi=\hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon^+_{\mathfrak{z}_{|\rho|}} u(z^{\rho}_{\pi_1})u(z^{\rho}_{\pi_2})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right],$$

and since the non-flat partitions in $\Pi[2 \times r]$ are made of pairs and singletons, this identity can be rewritten as the following consequence of Proposition 2.3, in which for simplicity of notation we write $\pi_1 = \{1, \ldots, r\}$ and $\pi_2 = \{r + 1, \ldots, 2r\}$.

Corollary 2.4 Assume that $u(z_1, \ldots, z_r; \omega) = 0$ whenever $z_i = z_j, 1 \le i \ne j \le r$, $\omega \in \Omega^X$. Then the second moment of the integral of k-processes is given by

$$\mathbb{E}\left[\left(\int_{X^r} u(z_1,\ldots,z_r;\omega)\omega(dz_1)\cdots\omega(dz_r)\right)^2\right]$$

= $\sum_{A\subset\pi_1} \frac{1}{(r-|A|)!} \sum_{\gamma:\pi_2\to A\cup\{r+1,\ldots,2r-|A|\}} \mathbb{E}\left[\int_{X^{2r-|A|}} \epsilon^+_{\mathfrak{z}_{2r-|A|}} u(z_{\pi_1})u(z_{\gamma(r+1)},\ldots,z_{\gamma(2r)})\hat{\lambda}^{2r-|A|}(d\mathfrak{z}_{2r-|A|})\right],$

where the above sum is over all bijections $\gamma: \pi_2 \to A \cup \{r+1, \ldots, 2r-|A|\}.$

Proof. We express the partitions $\rho \in \Pi[n \times r]$ with non-flat diagrams $\Gamma(\pi, \rho)$ in Proposition 3.1 as the collections of pairs and singletons

$$\rho = \{i, \gamma(i)\}\}_{i \in A} \cup \{\{i\}\}_{i \in \pi_1, i \notin A} \cup \{\{i\}\}_{i \in \pi_2, i \notin \gamma(A)},$$

for all subsets $A \subset \pi_1 = \{1, \ldots, r\}$ and bijections $\gamma : \pi_2 \to A \cup \{r+1, \ldots, 2r - |A|\}$.

In the case of 2-processes, Corollary 2.4 shows that

$$\mathbb{E}\left[\left(\int_{X^{2}} u(z_{1}, z_{2}; \omega)\omega(dz_{1})\omega(dz_{2})\right)^{2}\right] = \sum_{\substack{\rho \in \Pi[n \times 2] \\ \rho \wedge \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \epsilon_{\mathfrak{z}_{|\rho|}}^{+} \prod_{k=1}^{n} u(z_{\zeta^{\rho}(k,1)}, z_{\zeta^{\rho}(k,2)})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right]$$
$$= \sum_{\substack{A \subset \pi_{1} \\ \gamma: \{3,4\} \to A \cup \{3,\dots,4-|A|\}}} \frac{1}{(r-|A|)!} \mathbb{E}\left[\int_{X^{4-|A|}} \epsilon_{\mathfrak{z}_{4-|A|}}^{+} u(z_{1}, z_{2})u(z_{\gamma(3)}, z_{\gamma(4)})\hat{\lambda}^{4-|A|}(d\mathfrak{z}_{4-|A|})\right]$$

$$= \mathbb{E}\left[\int_{X^4} \epsilon_{\mathfrak{z}4}^+(u(z_1, z_2)u(z_3, z_4))\hat{\lambda}^4(d\mathfrak{z}_4)\right] \\ + \mathbb{E}\left[\int_{X^3} \epsilon_{\mathfrak{z}3}^+(u(z_1, z_2)u(z_1, z_3))\hat{\lambda}^3(d\mathfrak{z}_3)\right] + \mathbb{E}\left[\int_{X^3} \epsilon_{\mathfrak{z}3}^+(u(z_2, z_1)u(z_3, z_1))\hat{\lambda}^3(d\mathfrak{z}_3)\right] \\ + \mathbb{E}\left[\int_{X^3} \epsilon_{\mathfrak{z}3}^+(u(z_1, z_2)u(z_2, z_3))\hat{\lambda}^3(d\mathfrak{z}_3)\right] + \mathbb{E}\left[\int_{X^3} \epsilon_{\mathfrak{z}3}^+(u(z_2, z_1)u(z_3, z_2))\hat{\lambda}^3(d\mathfrak{z}_3)\right] \\ + \mathbb{E}\left[\int_{X^2} \epsilon_{\mathfrak{z}2}^+(u(z_1, z_2)u(z_1, z_2))\hat{\lambda}^2(d\mathfrak{z}_2)\right] + \mathbb{E}\left[\int_{X^2} \epsilon_{\mathfrak{z}2}^+(u(z_1, z_2)u(z_2, z_1))\hat{\lambda}^2(d\mathfrak{z}_2)\right].$$

Similarly, in the case of 3-processes we find

$$\begin{split} & \mathbb{E}\left[\left(\int_{X^3} u(z_1, z_2, z_3; \omega) \omega(dz_1) \omega(dz_2) \omega(dz_3)\right)^2\right] \\ &= \sum_{\substack{A \subseteq \{1, 2, 3\} \\ \gamma: \{4, 5, 6\} \to A \cup \{4, \dots, 6-|A|\}}} \frac{1}{(3 - |A|)!} \mathbb{E}\left[\int_{X^5} \epsilon^+_{45} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^5(d\mathfrak{z}_5)\right] \\ &= \mathbb{E}\left[\int_{X^6} \epsilon^+_{46} u(z_1, z_2, z_3) u(z_4, z_5, z_6) \hat{\lambda}^6(d\mathfrak{z}_6)\right] \\ &+ \frac{1}{2} \sum_{\gamma: \{4, 5, 6\} \to \{1, 5, 6\}} \mathbb{E}\left[\int_{X^5} \epsilon^+_{35} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^5(d\mathfrak{z}_5)\right] \\ &+ \frac{1}{2} \sum_{\gamma: \{4, 5, 6\} \to \{2, 5, 6\}} \mathbb{E}\left[\int_{X^5} \epsilon^+_{35} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^5(d\mathfrak{z}_5)\right] \\ &+ \frac{1}{2} \sum_{\gamma: \{4, 5, 6\} \to \{2, 5, 6\}} \mathbb{E}\left[\int_{X^5} \epsilon^+_{35} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^5(d\mathfrak{z}_5)\right] \\ &+ \frac{1}{2} \sum_{\gamma: \{4, 5, 6\} \to \{1, 2, 6\}} \mathbb{E}\left[\int_{X^4} \epsilon^+_{34} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^4(d\mathfrak{z}_4)\right] \\ &+ \sum_{\gamma: \{4, 5, 6\} \to \{1, 2, 3\}} \mathbb{E}\left[\int_{X^4} \epsilon^+_{34} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^4(d\mathfrak{z}_4)\right] \\ &+ \sum_{\gamma: \{4, 5, 6\} \to \{1, 2, 3\}} \mathbb{E}\left[\int_{X^4} \epsilon^+_{34} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^4(d\mathfrak{z}_4)\right] \\ &+ \sum_{\gamma: \{4, 5, 6\} \to \{1, 2, 3\}} \mathbb{E}\left[\int_{X^4} \epsilon^+_{34} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^4(d\mathfrak{z}_4)\right] \\ &+ \sum_{\gamma: \{4, 5, 6\} \to \{1, 2, 3\}} \mathbb{E}\left[\int_{X^4} \epsilon^+_{34} u(z_1, z_2, z_3) u(z_{\gamma(4)}, z_{\gamma(5)}, z_{\gamma(6)}) \hat{\lambda}^4(d\mathfrak{z}_4)\right] \end{aligned}$$

3 Random-connection model

Two point process vertices $x \neq y$ are independently connected in the random-connection graph with the probability H(x, y) given $\omega \in \Omega^X$, where $H : X \times X \longrightarrow [0, 1]$. In particular, the 1-hop count $\mathbb{1}_{\{x \leftrightarrow y\}}$ is a Bernoulli random variable with parameter H(x, y), and we have the relation

$$\mathbb{E}\left[\epsilon_{\mathfrak{z}_r}^+\prod_{i=0}^r\mathbb{1}_{\{z_i\leftrightarrow z_{i+1}\}}(\omega)\,\Big|\,\omega\right]=\prod_{i=0}^rH(z_i,z_{i+1})$$

for any subset $\{z_0, \ldots, z_{r+1}\}$ of distinct elements of X, where $\mathfrak{z}_r = \{z_1, \ldots, z_r\}$ and $x \leftrightarrow y$ means that $x \in X$ is connected to $y \in X$.

Given $x, y \in X$, the number of (r + 1)-hop sequences $z_1, \ldots, z_r \in \omega$ of vertices connecting x to y in the random graph is given by the multiparameter stochastic integral

$$N_{r+1}^{x,y} = \int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)$$

of the $\{0, 1\}$ -valued *r*-process

$$u(z_1, \dots, z_r; \omega) := \mathbb{1}_{\{z_i \neq z_j, \ 1 \le i < j \le r\}} \mathbb{1}_{\{z_1, \dots, z_r \in \omega\}} \prod_{i=0}^r \mathbb{1}_{\{z_i \leftrightarrow z_{i+1}\}}(\omega), \quad z_1, \dots, z_r \in X,$$
(3.1)

which vanishes on the diagonals in X^r , with $z_0 := x$ and $z_{r+1} := y$. In addition, for any distinct $z_1, \ldots, z_r \in X$ and $u(z_1, \ldots, z_r; \omega)$ given by (3.1) we have

$$\mathbb{E}\left[\epsilon_{\mathfrak{z}r}^+ u(z_1,\ldots,z_r;\omega) \mid \omega\right] = \mathbb{E}\left[\epsilon_{\mathfrak{z}r}^+ \prod_{i=0}^r \mathbb{1}_{\{z_i \leftrightarrow z_{i+1}\}}(\omega) \mid \omega\right] = \prod_{i=0}^r H(z_i,z_{i+1}), \quad (3.2)$$

therefore the first order moment of the (r + 1)-hop count between $x \in X$ and $y \in X$ is given as

$$\mathbb{E}\left[\int_{X^r} u(z_1, \dots, z_r; \omega) \omega(dz_1) \cdots \omega(dz_r)\right] = \mathbb{E}\left[\int_{X^r} \prod_{i=0}^r H(z_i, z_{i+1}) \hat{\lambda}^r(d\mathfrak{z}_r)\right], \quad (3.3)$$

see also Theorem II.1 of [9], as a consequence of the Georgii-Nguyen-Zessin identity (2.5).

In the next proposition we compute the moments of all orders of r-hop counts as sums over non-flat partition diagrams. The role of the powers $1/n_{l,i}^{\rho}$ in (3.4) is to ensure that all powers of H(x, y) in (3.4) are equal to one, since all powers of $\mathbb{1}_{\{z\leftrightarrow z'\}}$ in (3.5) below are equal to $\mathbb{1}_{\{z\leftrightarrow z'\}}$.

Proposition 3.1 The moment of order n of the (r + 1)-hop count between $x \in X$ and $y \in X$ is given by

$$\mathbb{E}\left[\left(N_{r+1}^{x,y}\right)^{n}\right] = \sum_{\substack{\rho \in \Pi[n \times r] \\ \rho \land \pi = \hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}} \prod_{l=1}^{n} \prod_{i=0}^{r} H^{1/n_{l,i}^{\rho}}(z_{\zeta^{\rho}(l,i)}, z_{\zeta^{\rho}(l,i+1)})\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right], \quad (3.4)$$

where $z_0 = x$, $z_{r+1} = y$, $\zeta^{\rho}(l, 0) = 0$, $\zeta^{\rho}(l, r+1) = r+1$, and

$$n_{l,i}^{\rho} := \#\{(p,j) \in \{1,\dots,n\} \times \{0,\dots,r\} : \{\zeta^{\rho}(l,i), \zeta^{\rho}(l,i+1)\} = \{\zeta^{\rho}(p,j), \zeta^{\rho}(p,j+1)\}\},\$$

$$1 \le l \le n, \ 0 \le i \le r.$$

Proof. Since $u(z_1, \ldots, z_r; \omega)$ vanishes whenever $z_i = z_j$ for some $1 \le i < j \le r$, by Proposition 2.3 we have

$$\mathbb{E}\left[\left(\int_{X^{r}} u(z_{1},\ldots,z_{r};\omega)\omega(dz_{1})\cdots\omega(dz_{r})\right)^{n}\right] \\
= \sum_{\substack{\rho\in\Pi[n\times r]\\\rho\wedge\pi=\hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}}\prod_{l=1}^{n}\prod_{i=0}^{r}\mathbb{1}_{\{z_{\zeta}\rho(l,i)\leftrightarrow z_{\zeta}\rho(l,i+1)\}}\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right] \\
= \sum_{\substack{\rho\in\Pi[n\times r]\\\rho\wedge\pi=\hat{0}}} \mathbb{E}\left[\int_{X^{|\rho|}}\prod_{l=1}^{n}\prod_{i=0}^{r}H^{1/n_{l,i}^{\rho}}(z_{\zeta}\rho(l,i),z_{\zeta}\rho(l,i+1))\hat{\lambda}^{|\rho|}(d\mathfrak{z}_{|\rho|})\right],$$
(3.5)

where we applied (3.2).

As in Corollary 2.4 we have the following consequence of Proposition 3.1, which is obtained by expressing the partitions $\rho \in \Pi[n \times r]$ with non-flat diagrams $\Gamma(\pi, \sigma)$ as a collection of pairs and singletons.

Corollary 3.2 The second moment of the (r+1)-hop count between $x \in X$ and $y \in X$ is given by

 $\mathbb{E}\left[\left(N_{r+1}^{x,y}\right)^2\right]$

$$= \sum_{\substack{A \subset \pi_1 \\ \gamma:\{1,\dots,r\} \to A \cup \{r+1,\dots,2r-|A|\}}} \frac{1}{(r-|A|)!} \mathbb{E}\left[\int_{X^{2r-|A|}} \prod_{i=0}^r H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^r H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{2r-|A|}(d\mathfrak{z}_{2r-|A|}) \right]$$

where the above sum is over all bijections $\gamma : \{1, \ldots, r\} \to A \cup \{r+1, \ldots, 2r - |A|\}$ with $\gamma(0) := 0$, $\gamma(r+1) =: r+1$, $z_0 =: x$, and $z_{r+1} := y$, and

$$n_{1,i}^{\gamma} = \#\{j \in \{0, \dots, r\} : \{i, i+1\} = \{\gamma(j), \gamma(j+1)\}\},\$$
$$n_{2,j}^{\gamma} = \#\{i \in \{0, \dots, r\} : (i, i+1) = (\gamma(j), \gamma(j+1))\},\$$

 $0 \leq i \leq r.$

Variance of 3-hop counts

When n = 2 and r = 2, Corollary 3.2 allows us to express the variance of the 3-hop count between $x \in X$ and $y \in X$ as follows:

$$\begin{aligned} &\operatorname{Var}\left[N_{3}^{x,y}\right] \\ &= \sum_{\substack{\emptyset \neq A \subset \{1,2\}\\ \gamma:\{1,2\} \to A \cup \{3,4-|A|\}}} \frac{1}{(2-|A|)!} \mathbb{E}\left[\int_{X^{4-|A|}} \prod_{i=0}^{2} H^{1/n_{1,i}^{\gamma}}(z_{i}, z_{i+1}) \prod_{j=0}^{2} H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{4-|A|}(d\mathfrak{z}_{4-|A|})\right] \\ &= \sum_{\substack{\gamma:\{1,2\} \to \{1,4\}}} \mathbb{E}\left[\int_{X^{3}} \prod_{i=0}^{2} H^{1/n_{1,i}^{\gamma}}(z_{i}, z_{i+1}) \prod_{j=0}^{2} H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{3}(dz_{1}, dz_{2}, dz_{4})\right] \\ &+ \sum_{\substack{\gamma:\{1,2\} \to \{2,4\}}} \mathbb{E}\left[\int_{X^{3}} \prod_{i=0}^{2} H^{1/n_{1,i}^{\gamma}}(z_{i}, z_{i+1}) \prod_{j=0}^{2} H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{3}(dz_{1}, dz_{2}, dz_{4})\right] \\ &+ \sum_{\substack{\gamma:\{1,2\} \to \{2,4\}}} \mathbb{E}\left[\int_{X^{2}} \prod_{i=0}^{2} H^{1/n_{1,i}^{\gamma}}(z_{i}, z_{i+1}) \prod_{j=0}^{2} H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^{2}(dz_{1}, dz_{2}, dz_{4})\right] \end{aligned}$$

Variance of 4-hop counts

When r = 3 and n = 2, Corollary 3.2 yields

$$\begin{split} &\operatorname{Var}\left[N_{4}^{x,y}\right] \\ &= \sum_{\substack{\emptyset \neq A \subset \pi_{1} \\ \gamma:\{1,\dots,3\} \to A \cup \{4,\dots,6-|A|\}}} \frac{1}{(3-|A|)!} \mathbb{E}\left[\int_{X^{6-|A|}} \prod_{i=0}^{3} H^{1/n_{1,i}^{\gamma}}(z_{i},z_{i+1}) \prod_{j=0}^{3} H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)},z_{\gamma(j+1)}) \hat{\lambda}^{6-|A|}(d\mathfrak{z}_{6-|A|})\right] \end{split}$$

$$\begin{split} &= \frac{1}{2} \sum_{\gamma:\{1,\dots,3\} \to \{1,5,6\}} \mathbb{E} \left[\int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5(dz_1, dz_2, dz_3, dz_5, dz_6) \right] \\ &+ \frac{1}{2} \sum_{\gamma:\{1,\dots,3\} \to \{2,5,6\}} \mathbb{E} \left[\int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5(dz_1, dz_2, dz_3, dz_5, dz_6) \right] \\ &+ \frac{1}{2} \sum_{\gamma:\{1,\dots,3\} \to \{3,5,6\}} \mathbb{E} \left[\int_{X^5} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^5(dz_1, dz_2, dz_3, dz_5, dz_6) \right] \\ &+ \sum_{\gamma:\{1,\dots,3\} \to \{1,2,6\}} \mathbb{E} \left[\int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\ &+ \sum_{\gamma:\{1,\dots,3\} \to \{1,3,6\}} \mathbb{E} \left[\int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\ &+ \sum_{\gamma:\{1,\dots,3\} \to \{2,3,6\}} \mathbb{E} \left[\int_{X^4} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^4(dz_1, dz_2, dz_3, dz_6) \right] \\ &+ \sum_{\gamma:\{1,\dots,3\} \to \{1,\dots,3\}} \mathbb{E} \left[\int_{X^3} \prod_{i=0}^3 H^{1/n_{1,i}^{\gamma}}(z_i, z_{i+1}) \prod_{j=0}^3 H^{1/n_{2,j}^{\gamma}}(z_{\gamma(j)}, z_{\gamma(j+1)}) \hat{\lambda}^3(dz_1, dz_2, dz_3, dz_6) \right] . \end{split}$$

4 Poisson case

In this section and the next one, we work in the Poisson random-connection model, using a Poisson point process on $X = \mathbb{R}^d$ with intensity $\lambda(dx)$ on \mathbb{R}^d . We let

$$H^{(n)}(x_0, x_n) := \int_{\mathbb{R}^d} \cdots \int_{\mathbb{R}^d} \prod_{i=0}^{n-1} H(x_i, x_{i+1}) \lambda(dx_1) \cdots \lambda(dx_{n-1}), \quad x_0, x_n \in \mathbb{R}^d, \quad n \ge 1.$$
(4.1)

The 2-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$ is given by the first order stochastic integral

$$\int_{\mathbf{R}^d} u(z;\omega)\omega(dz) = \int_{\mathbf{R}^d} \mathbb{1}_{\{x\leftrightarrow z_1\}} \mathbb{1}_{\{z_1\leftrightarrow y\}}(\omega)\omega(dz_1) = \int_{\mathbf{R}^d} \mathbb{1}_{\{x\leftrightarrow z_1\}} \mathbb{1}_{\{z_1\leftrightarrow y\}}\omega(dz_1),$$

and its moment of order n is

$$\mathbb{E}\left[\left(\int_{\mathbf{R}^{d}} u(z_{1};\omega)\omega(dz_{1})\right)^{n}\right] = \mathbb{E}\left[\left(\int_{\mathbf{R}^{d}} \mathbb{1}_{\{x\leftrightarrow z_{1}\}}\mathbb{1}_{\{z_{1}\leftrightarrow y\}}\omega(dz_{1})\right)^{n}\right]$$
$$= \sum_{\rho\in\Pi[n\times1]}\int_{X^{|\rho|}}\prod_{l=1}^{|\rho|} \left(H(x,z_{l})H(z_{l},y)\right)\lambda^{|\rho|}(dz_{1},\ldots,dz_{|\rho|})$$

$$= \sum_{k=1}^{n} S(n,k) \left(\int_{\mathbf{R}^d} H(x,z) H(z,y) \lambda(dz) \right)^k$$
$$= \sum_{k=1}^{n} S(n,k) \left(H^{(2)}(x,y) \right)^k,$$

therefore, from (2.1), the 2-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$ is a Poisson random variable with mean

$$H^{(2)}(x,y) = \int_{\mathbf{R}^d} H(x,z)H(z,y)\lambda(dz).$$

By (3.3), the first order moment of the *r*-hop count is given by

$$H^{(r)}(x,y) = \int_{X^{r-1}} \prod_{i=0}^{r-1} H(z_i, z_{i+1}) \lambda^{r-1}(dz_1, \dots dz_{r-1}).$$

Corollary 4.1 The variance of the r-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$ is given by

$$\operatorname{Var}\left[N_{r}^{x,y}\right] = \sum_{p=1}^{r-1} \sum_{\substack{1 \leq k_{1} < \dots < k_{p} < r \\ 1 \leq l_{1} < \dots < l_{p} < r}} \sum_{\sigma \in \Sigma[p]} \int_{X^{p}} \prod_{0 \leq i \leq p} H^{(k_{i+1}-k_{i})}(z_{i}, z_{i+1}) \prod_{\substack{0 \leq j \leq p \\ l_{\sigma(j+1)} - l_{\sigma(j)} + k_{j+1} - k_{j} > 2 \\ or \{j, j+1\} \neq \{\sigma(j), \sigma(j+1)\}}} H^{(l_{\sigma(j+1)} - l_{\sigma(j)})}(z_{\sigma(j)}, z_{\sigma(j+1)}) \lambda^{p}(d\mathfrak{z}_{p}),$$

with $k_0 = l_0 = 0$, $k_{p+1} = l_{p+1} = r$, $\sigma(0) = 0$, and $\sigma(r) = r$, where the above sum if over all permutations $\sigma \in \Sigma[p]$ of $\{1, \ldots, p\}$.

Proof. We rewrite the result of Corollary 3.2 by denoting the set $A \subset \pi_1$ as $A = \{k_1, \ldots, k_p\}$, for $1 \leq k_1 < \cdots < k_p \leq r-1$, and we identify $\gamma(A) \subset A \cup \{r+1, \ldots, 2r-|A|\}$ to $\{l_1, \ldots, l_p\}$, which requires a sum over the permutations of $\{1, \ldots, p\}$ since $1 \leq l_1 < \cdots < l_p \leq r-1$, where $1 \leq p \leq r-1$. In addition, the multiple integrals over contiguous index sets in A^c are evaluated using (4.1).

Variance of 3-hop counts

When n = 2 and r = 2 Corollary 4.1 allows us to compute the variance of the 3-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$, as follows:

$$\operatorname{Var}\left[N_{3}^{x,y}\right] \tag{4.2}$$

$$= 2 \int_{\mathbf{R}^d} H(x, z_1) H^{(2)}(z_1, y) H^{(2)}(z_1, y) \lambda(dz_1) + 2 \int_{\mathbf{R}^d} H(x, z_1) H^{(2)}(x, z_1) H^{(2)}(z_1, y) H(z_1, y) \lambda(dz_1) \\ + \int_{X^2} H(x, z_1) H(z_1, z_2) H(z_2, y) H(x, z_2) H(z_1, y) \lambda^2(dz_1, dz_2) + H^{(3)}(x, y).$$

By Corollary 4.1 the variance of 4-hop counts can be similarly computed explicitly as a sum of 33 terms, as follows:

$$\begin{aligned} \text{Var} \left[N_{4}^{x,y} \right] &= \int_{\mathbb{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta}^{(3)}(z_{1},y) H_{\beta}^{(3)}(z_{1},y) \lambda(dz_{1}) \end{aligned} \tag{4.3} \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta}^{(3)}(z_{1},y) H_{\beta}^{(2)}(x,z_{1}) H_{\beta}^{(2)}(z_{1},y) \lambda(dz_{1}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(x,z_{1})} H_{\beta}^{(3)}(x,z_{1}) H_{\beta}^{(3)}(z_{1},y) H_{\beta}(z_{1},y) \lambda(dz_{1}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(2)}(x,z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},y) \lambda(dz_{2}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(2)}(x,z_{2}) H_{\beta}^{(2)}(x,z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},y) \lambda(dz_{2}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(3)}(x,z_{3}) H_{\beta}(z_{3},y) H_{\beta}^{(3)}(x,z_{3}) H_{\beta}^{(2)}(z_{3},y) \lambda(dz_{3}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(3)}(x,z_{3}) H_{\beta}(z_{3},y) H_{\beta}^{(2)}(x,z_{3}) H_{\beta}^{(2)}(z_{3},y) \lambda(dz_{3}) \\ &+ \int_{\mathbb{R}^{d}} H_{\beta}^{(3)}(x,z_{3}) H_{\beta}(z_{3},y) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{1},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},z_{1}) H_{\beta}(z_{1},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(x_{2},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(z_{2},z_{1}) H_{\beta}(z_{1},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{2},y) H_{\beta}^{(2)}(x_{2},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{3},y) H_{\beta}^{(2)}(z_{3},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},z_{2}) H_{\beta}^{(2)}(z_{3},y) H_{\beta}^{(2)}(z_{3},y) \lambda^{2}(dz_{1},dz_{2}) \\ &+ \int_{X^{2}} H_{\beta}(x,z_{1}) H_{\beta}^{(2)}(z_{1},z_{3}) H_{\beta}^{(2)}(z_{3},y) H_{\beta}^{(2)}(z_{3},y) \lambda^{2}(dz_{1},dz_{3}) \end{aligned}$$

$$\begin{split} &+ \int_{X^2} H_{\beta}(x,z_1) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}(z_3,y) \lambda^2(dz_1,dz_3) \\ &+ \int_{X^2} H_{\beta}(x,z_1) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_3,z_1) H_{\beta}^{(2)}(z_1,y) \lambda^2(dz_1,dz_3) \\ &+ \int_{X^2} H_{\beta}(x,z_1) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}^{(2)}(x,z_1) H_{\beta}(z_1,z_3) H_{\beta}(z_3,y) \lambda^2(dz_1,dz_3) \\ &+ \int_{X^2} H_{\beta}(x,z_1) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}(z_3,y) H_{\beta}^{(2)}(x,z_3) H_{\beta}(z_3,y) \lambda^2(dz_1,dz_3) \\ &+ \int_{X^2} H_{\beta}(x,z_1) H_{\beta}^{(2)}(z_1,z_3) H_{\beta}(z_3,y) H_{\beta}^{(2)}(x,z_3) H_{\beta}(z_3,y) \lambda^2(dz_1,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}^{(2)}(x,z_3) H_{\beta}(z_3,y) \lambda^2(dz_2,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(x,z_2) H_{\beta}^{(2)}(z_2,z_3) H_{\beta}(z_3,y) \lambda^2(dz_2,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}^{(2)}(z_3,z_2) H_{\beta}(z_2,y) \lambda^2(dz_2,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}^{(2)}(z_3,z_2) H_{\beta}(z_2,y) \lambda^2(dz_2,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}^{(2)}(x,z_2) \lambda^2(dz_2,dz_3) \\ &+ \int_{X^2} H_{\beta}^{(2)}(x,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}^{(2)}(x,z_3) H_{\beta}(z_3,y) H_{\beta}(z_1,z_3) H_{\beta}(z_2,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_3,z_1) H_{\beta}(z_1,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_3,z_1) H_{\beta}(z_2,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_3,z_1) H_{\beta}(z_2,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_1,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_1,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}(x,z_3) H_{\beta}(z_1,y) \lambda^3(dz_1,dz_2,dz_3) \\ &+ \int_{X^3} H_{\beta}(x,z_1) H_{\beta}(z_1,z_2) H_{\beta}(z_2,z_3) H_{\beta}(z_3,y) H_{\beta}$$

5 Rayleigh fading

In this section we consider a Poisson point process on $X = \mathbb{R}^d$ with flat intensity $\lambda(dx) = \lambda dx$ on \mathbb{R}^d , $\lambda > 0$, and a Rayleigh fading function of the form

$$H_{\beta}(x,y) := e^{-\beta ||x-y||^2}, \qquad x, y \in \mathbb{R}^d, \quad \beta > 0.$$

Lemmas 5.1 and 5.2 can be used to evaluate the integrals appearing in Corollary 4.1and in the variance (4.2) of 3-hop counts.

Lemma 5.1 For all $n \ge 1, y_1, \ldots, y_n \in \mathbb{R}^d$ and $\beta_1, \ldots, \beta_n > 0$ we have

$$\int_{\mathbf{R}^d} \prod_{i=1}^n H_{\beta_i}(x, y_i) dx = \left(\frac{\pi}{\beta_1 + \dots + \beta_n}\right)^{d/2} \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left(y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i}\right).$$
Proof We start by showing that for all $n \ge 1$ we have

Proof. We start by showing that for all $n \ge 1$ we have

$$\prod_{i=1}^{n} H_{\beta_{i}}(x, y_{i})$$

$$= H_{\beta_{1}+\dots+\beta_{n}}\left(x, \frac{\beta_{1}y_{1}+\dots+\beta_{n}y_{n}}{\beta_{1}+\dots+\beta_{n}}\right) \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_{1}+\dots+\beta_{i})}{\beta_{1}+\dots+\beta_{i+1}}}\left(y_{i+1}, \frac{\beta_{1}y_{1}+\dots+\beta_{i}y_{i}}{\beta_{1}+\dots+\beta_{i}}\right).$$
(5.1)

Clearly, this relation holds for n = 1. In addition, at the rank n = 2 we have

$$\begin{aligned} H_{\beta_1}(x,y_1)H_{\beta_2}(x,y_2) &= e^{-\beta_1 \|y_1 - x\|^2} e^{-\beta_2 \|x - y_2\|^2} \\ &= e^{-\beta_1 \|y_1\|^2 - \beta_2 \|y_2\|^2 + 2\langle \beta_1 y_1 + \beta_2 y_2, x \rangle - (\beta_1 + \beta_2) \|x\|^2} \\ &= e^{-\beta_1 \|y_1\|^2 - \beta_2 \|y_2\|^2 - (\beta_1 + \beta_2) \|x - (\beta_1 y_1 + \beta_2 y_2)/(\beta_1 + \beta_2)\|^2 + \|\beta_1 y_1 + \beta_2 y_2\|^2/(\beta_1 + \beta_2)} \\ &= e^{-(\beta_1 + \beta_2) \|x - (\beta_1 y_1 + \beta_2 y_2)/(\beta_1 + \beta_2) \|^2 - \beta_1 \beta_2 \|y_1 - y_2\|^2/(\beta_1 + \beta_2)} \\ &= H_{\beta_1 + \beta_2} \left(x, \frac{\beta_1 y_1 + \beta_2 y_2}{\beta_1 + \beta_2} \right) H_{\frac{\beta_1 \beta_2}{\beta_1 + \beta_2}}(y_1, y_2), \end{aligned}$$

Next, assuming that (5.1) holds at the rank $n \ge 1$, we have

$$\prod_{i=1}^{n+1} H_{\beta_i}(x, y_i) = H_{\beta_{n+1}}(x, y_{n+1}) H_{\beta_1 + \dots + \beta_n} \left(x, \frac{\beta_1 y_1 + \dots + \beta_n y_n}{\beta_1 + \dots + \beta_n} \right)$$
$$\times \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left(y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i} \right)$$
$$= H_{\beta_1 + \dots + \beta_{n+1}} \left(x, \frac{\beta_1 y_1 + \dots + \beta_{n+1} y_{n+1}}{\beta_1 + \dots + \beta_n} \right) \prod_{i=1}^n H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left(y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i} \right).$$

As a consequence, we find

$$\int_{\mathbf{R}^d} \prod_{i=1}^n H_{\beta_i}(x, y_i) dx = \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left(y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i} \right)$$
$$\times \int_{\mathbf{R}^d} H_{\beta_1 + \dots + \beta_n} \left(x, \frac{\beta_1 y_1 + \dots + \beta_n y_n}{\beta_1 + \dots + \beta_n} \right) dx$$

$$= \left(\frac{\pi}{\beta_1 + \dots + \beta_n}\right)^{d/2} \prod_{i=1}^{n-1} H_{\frac{\beta_{i+1}(\beta_1 + \dots + \beta_i)}{\beta_1 + \dots + \beta_{i+1}}} \left(y_{i+1}, \frac{\beta_1 y_1 + \dots + \beta_i y_i}{\beta_1 + \dots + \beta_i}\right).$$

In particular, applying Lemma 5.1 for n = 2 yields

$$\int_{\mathbf{R}^{d}} H_{\beta_{1}}(y_{1}, x) H_{\beta_{2}}(x, y_{2}) dx = \left(\frac{\pi}{\beta_{1} + \beta_{2}}\right)^{d/2} H_{\frac{\beta_{1}\beta_{2}}{\beta_{1} + \beta_{2}}}(y_{1}, y_{2}) \qquad (5.2)$$

$$= \left(\frac{\pi}{\beta_{1} + \beta_{2}}\right)^{d/2} e^{-\beta_{1}\beta_{2} ||y_{1} - y_{2}||^{2}/(\beta_{1} + \beta_{2})}, \quad y_{1}, y_{2} \in \mathbf{R}^{d},$$

and the 2-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$ is a Poisson random variable with mean

$$H_{\beta}^{(2)}(x,y) = \lambda \int_{\mathbf{R}^{d}} H_{\beta}(x,z) H_{\beta}(z,y) dz$$
$$= \lambda \left(\frac{\pi}{2\beta}\right)^{d/2} H_{\beta/2}(x,y)$$
$$= \lambda \left(\frac{\pi}{2\beta}\right)^{d/2} e^{-\|x-y\|^{2}/2}.$$

By an induction argument similar to that of Lemma 5.1, we obtain the following lemma.

Lemma 5.2 For all $n \ge 1, x_0, \ldots, x_n \in \mathbb{R}^d$ and $\beta_1, \ldots, \beta_n > 0$ we have

$$\int_{\mathbf{R}^d} \cdots \int_{\mathbf{R}^d} \prod_{i=1}^n H_{\beta_i}(x_{i-1}, x_i) dx_1 \cdots dx_{n-1}$$
$$= \left(\frac{\pi^{n-1}}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n} \right)^{d/2} H_{\frac{\beta_1 \cdots \beta_n}{\sum_{i=1}^n \beta_1 \cdots \beta_{i-1} \beta_{i+1} \cdots \beta_n}}(x_0, y_n).$$

Proof. Clearly, the relation holds at the rank n = 1. Assuming that it holds at the rank $n \ge 1$ and using (5.2), we have

$$\int_{\mathbf{R}^d} \cdots \int_{\mathbf{R}^d} \prod_{i=1}^{n+1} H_{\beta_i}(x_{i-1}, x_i) dx_1 \cdots dx_n$$
$$= \int_{\mathbf{R}^d} H_{\beta_{n+1}}(x_n, x_{n+1}) \int_{\mathbf{R}^d} \cdots \int_{\mathbf{R}^d} \prod_{i=1}^n H_{\beta_i}(x_{i-1}, x_i) dx_1 \cdots dx_n$$

$$= \left(\frac{\pi^{n-1}}{\sum_{i=1}^{n}\beta_{1}\cdots\beta_{i-1}\beta_{i+1}\cdots\beta_{n}}\right)^{d/2} \int_{\mathbf{R}^{d}} H_{\frac{\beta_{1}\cdots\beta_{n}}{\sum_{i=1}^{n}\beta_{1}\cdots\beta_{i-1}\beta_{i+1}\cdots\beta_{n}}}(x_{0},x_{n})H_{\beta_{n+1}}(x_{n},x_{n+1})dx_{n}$$

$$= \left(\frac{\pi^{n-1}}{\sum_{i=1}^{n}\beta_{1}\cdots\beta_{i-1}\beta_{i+1}\cdots\beta_{n}}\right)^{d/2} \left(\frac{\pi}{\frac{\beta_{1}\cdots\beta_{n}}{\sum_{i=1}^{n}\beta_{1}\cdots\beta_{i-1}\beta_{i+1}\cdots\beta_{n}}} + \beta_{n+1}\right)^{d/2} H_{\frac{\beta_{1}\cdots\beta_{n+1}}{\sum_{i=1}^{n+1}\beta_{1}\cdots\beta_{i-1}\beta_{i+1}\cdots\beta_{n+1}}}(x_{0},x_{n+1}).$$

In particular, the first order moment of the *r*-hop count between $x_0 \in \mathbb{R}^d$ and $x_r \in \mathbb{R}^d$ is given by

$$H_{\beta}^{(r)}(x_{0}, x_{r}) = \int_{\mathbb{R}^{d}} \cdots \int_{\mathbb{R}^{d}} \prod_{i=0}^{r-1} H_{\beta}(x_{i}, x_{i+1}) \lambda(dx_{1}) \cdots \lambda(dx_{r-1})$$

$$= \lambda^{r-1} \left(\frac{\pi^{r-1}}{r\beta^{r-1}}\right)^{d/2} H_{\beta/r}(x, y)$$

$$= \lambda^{r-1} \left(\frac{\pi^{r-1}}{r\beta^{r-1}}\right)^{d/2} e^{-\beta ||x-y||^{2}/r}, \quad x, y \in \mathbb{R}^{d}.$$
(5.3)

Variance of 3-hop counts

Corollary 4.1 and Lemma 5.2 allow us to recover Theorem II.3 of [9], for the variance of 3-hop counts by a shorter argument, while extending it from the plane $X = \mathbb{R}^2$ to $X = \mathbb{R}^d$.

Corollary 5.3 The variance of the 3-hop count between $x \in \mathbb{R}^d$ and $y \in \mathbb{R}^d$ is given by

$$\operatorname{Var}\left[N_{3}^{x,y}\right] = 2\lambda^{3} \left(\frac{\pi^{3}}{8\beta^{3}}\right)^{d/2} e^{-\beta \|x-y\|^{2}/2} + \lambda^{2} \left(\frac{\pi^{2}}{3\beta^{2}}\right)^{d/2} e^{-\beta \|x-y\|^{2}/3} \\ + 2\lambda^{3} \left(\frac{\pi^{3}}{12\beta^{3}}\right)^{d/2} e^{-3\beta \|x-y\|^{2}/4} + \lambda^{2} \left(\frac{\pi^{2}}{8\beta^{2}}\right)^{d/2} e^{-\beta \|x-y\|^{2}}.$$

Proof. By (5.3) and Lemma 5.2 we have

$$\begin{split} &\int_{\mathbf{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta}^{(2)}(z_{1},y) H_{\beta}^{(2)}(z_{1},y) \lambda(dz_{1}) = \lambda^{2} \left(\frac{\pi^{2}}{4\beta^{2}}\right)^{d/2} \int_{\mathbf{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta/2}^{2}(z_{1},y) \lambda(dz_{1}) \\ &= \lambda^{3} \left(\frac{\pi^{2}}{4\beta^{2}}\right)^{d/2} \int_{\mathbf{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta}(z_{1},y) \lambda(dz_{1}) = \lambda^{3} \left(\frac{\pi^{3}}{8\beta^{3}}\right)^{d/2} H_{\beta/2}(x,y), \\ &\int_{\mathbf{R}^{d}} H_{\beta}(x,z_{1}) H_{\beta}^{(2)}(x,z_{1}) H_{\beta}^{(2)}(z_{1},y) H_{\beta}(z_{1},y) \lambda(dz_{1}) \end{split}$$

$$\begin{split} &= \lambda^2 \left(\frac{\pi^2}{4\beta^2}\right)^{d/2} \int_{\mathbf{R}^d} H_{3\beta/2}(z_1, y) H_{3\beta/2}(x, z_1) \lambda(dz_1) = \lambda^3 \left(\frac{\pi^3}{12\beta^3}\right)^{d/2} H_{3\beta/4}(x, y), \\ &\int_{X^2} H_{\beta}(x, z_1) H_{\beta}(z_1, z_2) H_{\beta}(z_2, y) H_{\beta}(x, z_2) H_{\beta}(z_1, y) \lambda^2(dz_1, dz_2) \\ &= \lambda \left(\frac{\pi}{3\beta}\right)^{d/2} H_{\beta}(x, y) \int_{\mathbf{R}^d} H_{2\beta/3}(z_2, (x+y)/2) H_{2\beta}(z_2, (x+y)/2) \lambda(dz_2) \\ &= \lambda^2 \left(\frac{\pi^2}{8\beta^2}\right)^{d/2} H_{\beta}(x, y), \end{split}$$

and we conclude by (4.2).

References

- B. Bassan and E. Bona. Moments of stochastic processes governed by Poisson random measures. Comment. Math. Univ. Carolin., 31(2):337–343, 1990.
- [2] K. Bogdan, J. Rosiński, G. Serafin, and L. Wojciechowski. Lévy systems and moment formulas for mixed Poisson integrals. In *Stochastic analysis and related topics*, volume 72 of *Progr. Probab.*, pages 139–164. Birkhäuser/Springer, Cham, 2017.
- [3] K.N. Boyadzhiev. Exponential polynomials, Stirling numbers, and evaluation of some gamma integrals. *Abstr. Appl. Anal.*, Art. ID 168672:18 pages, 2009.
- [4] J.-C. Breton and N. Privault. Factorial moments of point processes. Stochastic Processes and their Applications, 124(10):3412–3428, 2014.
- [5] L. Decreusefond and I. Flint. Moment formulae for general point processes. J. Funct. Anal., 267:452–476, 2014.
- [6] L. Decreusefond, I. Flint, N. Privault, and G.L Torrisi. Determinantal point processes. In G. Peccati and M. Reitzner, editors, *Stochastic Analysis for Poisson Point Processes: Malliavin Calculus, Wiener-Itô Chaos Expansions and Stochastic Geometry*, volume 7 of *Bocconi & Springer Series*, pages 311–342, Berlin, 2016. Springer.
- [7] N. Deng, W. Zhou, and M. Haenggi. The Ginibre point process as a model for wireless networks with repulsion. *IEEE Transactions on Wireless Communications*, 14:107–121, 2015.
- [8] O. Kallenberg. Random measures. Akademie-Verlag, Berlin, fourth edition, 1986.
- [9] A.P. Kartun-Giles and S. Kim. Counting k-hop paths in the random connection model. IEEE Transactions on Wireless Communications, 17(5):3201–3210, 2018.
- [10] H. B. Kong, I. Flint, P. Wang, D. Niyato, and N. Privault. Exact performance analysis of ambient RF energy harvesting wireless sensor networks with Ginibre point process. *IEEE Journal on Selected Areas in Communications*, 34:3769–3784, 2016.
- [11] J. Mecke. Stationäre zufällige Masse auf lokalkompakten Abelschen Gruppen. Z. Wahrscheinlichkeitstheorie Verw. Geb., 9:36–58, 1967.
- [12] R. Meester and R. Roy. Continuum percolation, volume 119 of Cambridge Tracts in Mathematics. Cambridge University Press, Cambridge, 1996.

- [13] N. Miyoshi and T. Shirai. A cellular network model with Ginibre configurated base stations. Research Rep. on Math. and Comp. Sciences (Tokyo Inst. of Tech.), 2012.
- [14] X.X. Nguyen and H. Zessin. Integral and differential characterization of the Gibbs process. Math. Nachr., 88:105–115, 1979.
- [15] G. Peccati and M. Taqqu. Wiener Chaos: Moments, Cumulants and Diagrams: A survey with Computer Implementation. Bocconi & Springer Series. Springer, 2011.
- [16] N. Privault. Moment identities for Poisson-Skorohod integrals and application to measure invariance. C. R. Math. Acad. Sci. Paris, 347:1071–1074, 2009.
- [17] N. Privault. Invariance of Poisson measures under random transformations. Ann. Inst. H. Poincaré Probab. Statist., 48(4):947–972, 2012.
- [18] N. Privault. Moments of Poisson stochastic integrals with random integrands. Probability and Mathematical Statistics, 32(2):227–239, 2012.
- [19] N. Privault. Combinatorics of Poisson stochastic integrals with random integrands. In G. Peccati and M. Reitzner, editors, Stochastic Analysis for Poisson Point Processes: Malliavin Calculus, Wiener-Itô Chaos Expansions and Stochastic Geometry, volume 7 of Bocconi & Springer Series, pages 37–80. Springer, Berlin, 2016.
- [20] I.M. Slivnyak. Some properties of stationary flows of homogeneous random events. Theory Probab. Appl., 7(3):336–341, 1962.